

# Associations of Combined Exposures to Surrounding Green, Air Pollution, and Road Traffic Noise with Cardiometabolic Diseases

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**BACKGROUND:** Surrounding green, air pollution, and noise have been associated with cardiometabolic diseases, but most studies have assessed only one of these correlated exposures.

**OBJECTIVES:** We aimed to evaluate associations of combined exposures to green, air pollution, and road traffic noise with cardiometabolic diseases.

**METHODS:** In this cross-sectional study, we studied associations between self-reported physician-diagnosed diabetes, hypertension, heart attack, and stroke from a Dutch national health survey of 387,195 adults and residential surrounding green, annual average air pollutant concentrations [including particulate matter with aerodynamic diameter  $\leq 10 \mu\text{m}$  ( $\text{PM}_{10}$ ),  $\text{PM}$  with aerodynamic diameter  $\leq 2.5 \mu\text{m}$  ( $\text{PM}_{2.5}$ ), nitrogen dioxide ( $\text{NO}_2$ ), and oxidative potential (OP) with the dithiothreitol (DTT) assay ( $\text{OP}^{\text{DTT}}$ )] and road traffic noise. Logistic regression models were used to analyze confounding and interaction of surrounding green, air pollution, and noise exposure.

**RESULTS:** In single-exposure models, surrounding green was inversely associated with diabetes, while air pollutants ( $\text{NO}_2$ ,  $\text{OP}^{\text{DTT}}$ ) and road traffic noise were positively associated with diabetes. In two-exposure analyses, associations with green and air pollution were attenuated but remained. The association between road traffic noise and diabetes was reduced to unity when adjusted for surrounding green or air pollution. Air pollution and surrounding green, but not road traffic noise, were associated with hypertension in single-exposure models. The weak inverse association of surrounding green with hypertension attenuated and lost significance when adjusted for air pollution. Only  $\text{PM}_{2.5}$  was associated with stroke and heart attack.

**CONCLUSIONS:** Studies including only one of the correlated exposures surrounding green, air pollution, and road traffic noise may overestimate the association of diabetes and hypertension attributed to the studied exposure. <https://doi.org/10.1289/EHP3857>

## Introduction

In daily life, humans are exposed to a mixture of environmental exposures that can affect health. Exposure to air pollution and road traffic noise, for example, have both been associated with increased risks of diabetes and hypertension (Wang et al. 2014a; Balti et al. 2014; Eze et al. 2015; Dzhambov 2015; Fuks et al. 2014; Van Kempen and Babisch 2012). Air pollution and road traffic noise may induce inflammation, oxidative stress, vascular dysfunction, and autonomic nervous system imbalance, thereby contributing to the development of cardiometabolic diseases (Münzel et al. 2016b). On the other hand, research suggests that exposure to surrounding green is associated with reduced risks of cardiometabolic disease (Dalton et al. 2016; Astell-Burt, et al. 2014; Ngom et al. 2016; Maas et al. 2009). Surrounding green has been hypothesized to benefit health via several pathways, including promoting physical activity and social cohesion and decreasing psychological stress (Hartig et al. 2014; Nieuwenhuijsen et al. 2017).

Long-term exposures to surrounding green, air pollution, and road traffic noise are, to a certain extent, related to each other. Higher levels of surrounding green, for example, tend to go together with lower levels of air pollution and traffic noise because of absence of air pollution and noise sources in green areas or

removal of air pollutants from the air (Markevych et al. 2017). Road traffic noise is related to traffic intensity, which also affects traffic related air pollutants such as nitrogen dioxide ( $\text{NO}_2$ ) and black carbon [ $\text{PM}_{2.5\text{absorbance}}$  ( $\text{PM}_{2.5\text{abs}}$ )], which can lead to spatial correlation between air pollution and road traffic noise (Münzel et al. 2016a; Têtreault et al. 2013). Correlations between surrounding green, air pollution, and road traffic noise depend on the study setting but are typically moderate (Boogaard et al. 2009; Fecht et al. 2016; Hystad et al. 2014; Thiering et al. 2016; Davdand et al. 2012).

The large majority of studies on the associations of long-term exposure to surrounding green, air pollution, or road traffic noise have evaluated single-exposure associations of one of these environmental exposures, ignoring potential confounding by and interaction with the other two environmental exposures. Therefore, it is difficult to interpret the results of studies that only evaluated single exposures. Only a few studies investigated the combined effects of these environmental exposures with cardiometabolic diseases. Some of these studies suggest that exposure to air pollution confounds associations with exposure to road traffic noise and vice versa (Balti et al. 2014; Eze et al. 2017; Floud et al. 2013; Bodin et al. 2016; Selander et al. 2009; Sørensen et al. 2013), while there are also indications for interaction effects between air pollution and road traffic noise (Sørensen et al. 2014).

Studies that evaluated associations of surrounding green with health have treated air pollution and traffic noise as confounders (Hystad et al. 2014; Crouse et al. 2017; Thiering et al. 2016; Villeneuve et al. 2012) or as mediators (James et al. 2016; Gascon et al. 2018; Dzhambov et al. 2018). The difference between treating them as confounders or mediator is the assumption of causality of the relationships of surrounding green with air pollution and road traffic noise. If air pollution and road traffic noise are on the causal pathway from surrounding green to health, mediation analysis is appropriate. Such a causal relationship is supported by the fact that more surrounding green may lead to lower road traffic noise and air pollution concentrations by limiting dispersion of noise and air pollution through green barriers or by scavenging of air pollution. These mechanisms, however, explain only part of the empirical

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correlations of surrounding green with air pollution and road traffic noise. More important is the fact that in a greener area, there are fewer sources of air pollution and road traffic noise and, consequently, lower air pollution and noise levels. This reflects a common source (or lack thereof) and thus not a causal pathway from surrounding green to health. Hence, air pollution and road traffic noise can be considered both as confounders and as mediators in the analysis of the relationship between surrounding green and health.

Strak et al. (2017) previously reported associations between exposure to air pollution and diabetes in a large national health survey of 387,195 citizens aged  $\geq 19$  y. For the current study, surrounding green and road traffic noise were also added to the health survey data. Further, besides diabetes, we also included hypertension, stroke, and heart attack. For all these health end points, there is some evidence for associations with surrounding green, air pollution, and road traffic noise from typically single-exposure studies. The aim of this study was to evaluate associations of combined long-term exposure to surrounding green, air pollution, and road traffic noise with prevalence of diabetes, hypertension, stroke, and heart attack.

## Methods

We specifically analyzed potential confounding and interaction of long-term exposure to surrounding green, air pollution, and road traffic noise. In addition, we analyzed mediation to evaluate whether decreased levels of air pollution and road traffic noise are possible mechanisms underlying potential beneficial associations of surrounding green with cardiometabolic outcomes. For the mediation analyses, we thus assumed that air pollution and road traffic noise would be on the causal pathway from surrounding green to health, while for the confounding analyses, we assumed that air pollution and road traffic noise were not on the causal pathway from surrounding green to health. Figure 1 provides a graphical representation of the four approaches (single exposure, confounding, interaction, and mediation). Air pollution and road traffic noise were treated as mutual confounders, because one exposure is not on the causal pathway of the other exposure to health.

## Study Design and Study Population

We used cross-sectional data from a national health survey [Dutch Public Health Monitor 2012 (PHM) (*Gezondheidsmonitor Volwassenen GGD-en, CBS en RIVM*)] that covered issues related to personal characteristics, lifestyle, socioeconomic status (SES), and physical and mental health. The PHM was conducted by 28 public health services (GGD-en), Statistics Netherlands (CBS), and the National Institute for Public Health and the Environment (RIVM) in 2012. The elderly ( $\geq 65$  y) were oversampled as part of the design of the PHM. The PHM includes information on 387,195 citizens aged  $\geq 19$  y; the response rate was 45–50%. The home address of each subject was geocoded; hence, no individuals were excluded due to missing geocodes.

Statistics Netherlands enriched the PHM with information on standardized household income and country of origin. Standardized household income was adjusted for differences in household size and composition. The PHM was also linked with information on SES at the neighborhood level (four-digit postal code area). This indicator represents the educational, occupational, and economical status of the neighborhood and is derived by the Netherlands Institute for Social Research (Knol 1998). A high SES neighborhood score indicates a high SES of the neighborhood. We used the 2002 SES score. The 2002 SES correlated highly with the 2010 SES score, which became available in a late phase of the study ( $r_s = 0.86$ ).

## Outcome Definition

The PHM included questions on diabetes, hypertension, stroke, and heart attack. The main questions were: “Do you have diabetes?”, “Do you have hypertension?”, “Have you ever had a stroke?” and “Have you ever had a heart attack?”. These questions were followed by a question “In the last 12 months, have you been treated for ... by a physician or a specialist?”. Strak et al. (2017) previously showed that 92% of positive answers to the main question for diabetes were also positive in the question about treatment by a physician or specialist in this study population. Moreover, nearly 85% of the population with self-reported physician-diagnosed diabetes in the past 12 months have also been prescribed diabetes medication, based on an external database (Strak et al. 2017). For hypertension, 91% of the positive answers for the main question were also positive in the question about treatment by a physician or specialist. For heart attack, 70% of the people who ever had a heart attack reported that they have been treated by a physician or a specialist within the last 12 months. For stroke, this was 49%.

Self-reports from questionnaires have been found to agree moderately to very well with medical records for diabetes, hypertension, and heart attack; kappa coefficients ranged from 0.76 to 0.80 for diabetes, from 0.56 to 0.75 for hypertension, and from 0.48 to 0.80 for heart attack (Okura et al. 2004; Schneider et al. 2012; Hansen et al. 2014; Muggah et al. 2013; Machón et al. 2013). Agreements for stroke varied per study; kappa statistics ranged from 0.35 to 0.71 (Okura et al. 2004; Schneider et al. 2012; Muggah et al. 2013; Machón et al. 2013; Hansen et al. 2014).

We used the questions about treatment by a physician or specialist to define health outcomes. As agreements between the main question and the question about treatment were moderate for heart attack and stroke, the first question (“Have you ever had a heart attack/stroke”) was used as the health outcome in sensitivity analyses.

## Exposure Assessment

**Surrounding green.** We used two green metrics to assess exposure to surrounding green that have been described previously (Klompaker et al. 2017). In brief, the Normalized Difference Vegetation Index (NDVI) was used to assess surrounding greenness, i.e., average density of green vegetation within a circular buffer around the participant’s residential address. The NDVI was derived from LANDSAT 5 TM and captures the density of green vegetation at a spatial resolution of  $30 \times 30$  m based on land surface reflectance of visible (red) and near-infrared parts of spectrum (<https://earthobservatory.nasa.gov>). NDVI values range between  $-1$  and  $1$ , with higher numbers indicating more photosynthetically active greenness and thus a higher density of green vegetation. Negative NDVI values were set to zero (Klompaker et al. 2017). We combined cloud-free images from the summer of 2010 to create a map that covers (almost) the whole country.

A highly detailed national land-use database of the Netherlands of 2010 [TOP10NL, CC-BY Kadaster, <https://www.kadaster.com/automatic-generalisation> (Kadaster 2010)] was used to assess surrounding green space, i.e., the proportion of green space within a buffer around the participant’s residential address. TOP10NL divides the Netherlands into polygons with different classes of land use (water, road, and terrain). The terrain class is divided in 21 sub-classes; 11 of these classes correspond to green areas. TOP10NL does, in contrast to the NDVI, not include private green property (such as gardens) and street greenery.

Surrounding greenness and surrounding green space were assessed in buffers with radii of 100; 300; 500; 1,000; and 3,000 m for all addresses in the Netherlands. Analyses to assess surrounding green were performed in ArcGIS (version 10.2.2; Esri).

**Air pollution.** For each home address, long-term average concentrations of particulate matter [PM with aerodynamic diameter  $\leq 10\ \mu\text{m}$  ( $\text{PM}_{10}$ ),  $\text{PM}_{\text{coarse}}$ , PM with aerodynamic diameter  $\leq 2.5\ \mu\text{m}$  ( $\text{PM}_{2.5}$ ),  $\text{PM}_{2.5\text{abs}}$ ] and  $\text{NO}_2$  were assessed by land-use regression (LUR) models developed within the framework of the European Study of Cohorts for Air Pollution Effects (ESCAPE) project (Beelen et al. 2013; Eeftens et al. 2012).  $\text{NO}_2$  levels higher than  $80\ \mu\text{g}/\text{m}^3$  ( $n = 18$ ) were set to  $80\ \mu\text{g}/\text{m}^3$ , as these values are probably due to an unrealistic combination of explanatory variables. Further, we used long-term average concentrations of two oxidative potential (OP) metrics: electron spin resonance ( $\text{OP}^{\text{ESR}}$ ) and dithiothreitol ( $\text{OP}^{\text{DTT}}$ ) (Yang et al. 2015). OP is an intrinsic measure of PM to oxidize target molecules, and thus, it effectively incorporates biologically relevant properties of PM (Yang et al. 2015). Performances of LUR models were evaluated using leave-one-out cross validation ( $\text{R}^2_{\text{LOOCV}}$ ) and ranged from 0.38 for  $\text{PM}_{\text{coarse}}$  to 0.89 for  $\text{PM}_{2.5\text{abs}}$  (Table S1). Models were based on air pollution measurements conducted in 2009, 3 y prior to the health survey. There were no LUR models available for 2012. Previous studies have shown that the spatial variation of air pollution and traffic noise exposure levels remain stable over periods of about 10 y in Western countries (Eeftens et al. 2011; Fecht et al. 2016).

**Road traffic noise.** Residential road traffic noise levels were assessed by the Standard Model Instrumentation for Noise Assessments (STAMINA). STAMINA is a model to map environmental noise in the Netherlands. This model was developed at the Dutch National Institute for Public Health and the Environment (RIVM) and uses the standard Dutch calculation method for traffic and industrial noise (Schreurs et al. 2010). The spatial resolution of the noise maps depends on the distance between source and observation point. The lowest resolution is  $80 \times 80\ \text{m}$ , and close to the source the level of detail is the highest, with a resolution of  $10 \times 10\ \text{m}$  (Schreurs et al. 2010). This method is used in the Netherlands to implement the European Environmental Noise Directive (EU Directive 2002).

Daily average (24 h,  $\text{L}_{\text{den}}$ ) and nighttime average (2300–0700 hours,  $\text{L}_{\text{night}}$ ) road traffic noise exposures were assessed for 2011. Because  $\text{L}_{\text{den}}$  and  $\text{L}_{\text{night}}$  were highly correlated [Spearman's rank order correlation coefficient ( $r_s$ ) = 0.99], we only used  $\text{L}_{\text{den}}$  in our analyses.

### Exclusion of Subjects

Since we did not have land-use data across the border of the Netherlands, subjects with residential addresses within 3 km (largest buffer) of the border of the Netherlands were excluded from our analyses (8.4%).

### Statistical Analyses

We calculated Spearman correlations to evaluate relations between the exposure variables at PHM subject addresses. We developed logistic regression models to study whether the exposure variables were associated with cardiometabolic diseases (diabetes, hypertension, stroke, and heart attack). We performed single-exposure, multi-exposure (confounding and interaction), and mediation analyses. An example of the different analyses performed for surrounding greenness (NDVI 300 m), air pollution ( $\text{OP}^{\text{DTT}}$ ), and diabetes is shown in Figure 1. Further, we calculated Spearman correlations between surrounding green, road traffic noise, and measured air pollution at the Dutch ESCAPE measurement sites to evaluate whether correlations between modeled and measured exposures differed.

All analyses were performed with R (version 3.3.1; R Development Core Team).

**Single-exposure regression models.** We used single-exposure regression models to analyze individual associations of exposures. We specified *a priori* several regression models with increasing degree of covariate adjustment. Model 1 included age and sex. Model 2 was additionally adjusted for individual SES variables (marital status, country of origin, work, standardized household income, level of education). Model 3 further included smoking habits (current, former, never), number of cigarettes/day, alcohol consumption (current, former, never), number of alcohol glasses/week, physical activity, and body mass index. Model 4 was additionally adjusted for area-level SES (main model). Categories of covariates in the regression analyses were identical to the categories presented in Table 1, except for age (12 categories: 19–24, 25–29, 30–34, 35–39, 40–44, 45–49, 50–54, 55–59, 60–64, 65–74, 75–84,  $\geq 85\ \text{y}$ ).

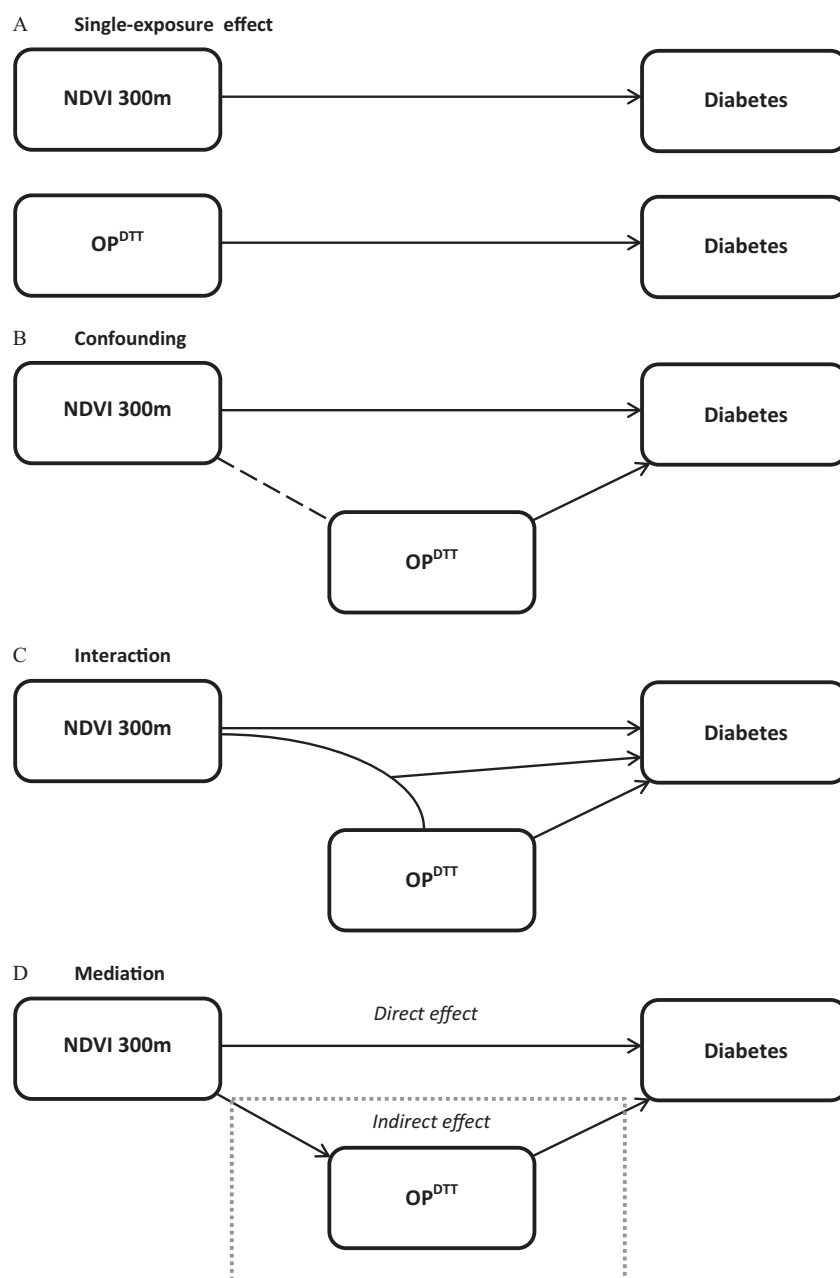
For each exposure variable, we specified a regression model with the exposure variable as linear term [1 degree of freedom (df)]. Additionally, we used natural splines (3 df) to determine the linearity of exposure–response relationships (only for the main model). To test whether the goodness of fit of the models with splines was significantly better than the goodness of fit of the linear models, we used the likelihood ratio test and defined statistical significance by a two-sided alpha level of  $p = 0.05$ . Exposure–response curves showed a nonsignificant or small to moderate deviation from linearity for most associations (Figure S1). Deviations from linearity primarily occurred in areas with very few data, as evidenced by the wide confidence intervals (CIs) of the nonlinear part of most curves. We considered the deviations from linearity to be sufficiently large to present both linear and quintile analyses in the paper, but not large enough to make the linear analysis uninformative. We *a priori* decided to use quintiles instead of tertiles or quartiles to obtain a finer assessment, which was made possible by the large study population. We presented the linear effect estimate of exposure to surrounding green and air pollution per interquartile range (IQR) and for road traffic noise per 5 dB. We used 5 dB instead of the IQR, as this is the standard for reporting of noise effect estimates.

Surrounding green in different buffer sizes was highly correlated, and associations between closely related buffer sizes, such as 1,000 and 3,000 m, barely differed (data not shown). Hence, we decided to only report associations of the 300- and 1,000-m buffers and focus on these buffers in the multi-exposure models to limit the number of reported analyses.

As sensitivity analyses, we evaluated associations of surrounding green, air pollution, and road traffic noise with cardiometabolic diseases in the elderly population ( $\geq 75\ \text{y}$ ,  $n = 62,111$ ). Further, we evaluated the associations of exposures with heart attack and stroke by the questions if participants ever had a heart attack ( $n = 14,675$ ) or stroke ( $n = 12,068$ ) instead of if they have been treated by a physician or a specialist within the last 12 months for a heart attack or stroke.

**Multi-exposure regression models.** We used multi-exposure models to analyze whether a potential association of one of the environmental exposures with cardiometabolic health was confounded by other environmental exposures. Here, we specified surrounding green, air pollution, and road traffic noise as mutual confounders in the analyses with cardiometabolic health and thus assumed a noncausal relationship between surrounding green and air pollution and road traffic noise.

To evaluate potential mutual confounding of exposures, we specified two-exposure models. We performed multi-exposure regression analyses with combinations of surrounding green, air pollutants, and road traffic noise exposures. We used generalized variance inflation factors to quantify multicollinearity between the exposures (Fox and Monette 1992).



**Figure 1.** Schematic overview of the analyses performed for two exposures and one health outcome. The black solid lines represent hypothesized causal relationships; the black dashed line represents a noncausal relationship. Note: NDVI, Normalized Difference Vegetation Index; OP<sup>DTT</sup>, oxidative potential (OP) metric with dithiothreitol (DTT).

**Joint odds ratio.** In epidemiological studies, it can be difficult to disentangle individual effects of correlated exposures, as exposure estimates share overlapping information. Hence, in exposure–health models, the estimated coefficient for one exposure based on a single-exposure model could be partly or fully explained (confounded) by the overlapping information with another exposure, depending on the extent of overlap and the associations with the outcome. Therefore, the estimated coefficient for an exposure from a single-exposure model could be an overestimation of the true coefficient of that exposure. However, because no adjustment for the overlapping information with the other exposure is made in a single-exposure model, the coefficient of the single-exposure model could be a good indication for the joint coefficient of both exposures.

A joint coefficient of co-occurring exposures can be assessed using the cumulative risk index (CRI) method (Crouse et al.

2015; Jerrett et al. 2013; Lippmann et al. 2013). Previous studies used the CRI to find out which combinations of air pollutants are required to fully capture the toxicity of the mix of air pollutants (Jerrett et al. 2013; Crouse et al. 2015). We used the CRI to investigate which combination of exposures is needed to assess the joint impact of surrounding green, air pollution, and road traffic noise on cardiometabolic diseases. As we found negative correlations between surrounding green, air pollution, and road traffic noise, exposure to a combination of low levels of surrounding green and high levels of air pollution and road traffic noise or the other way around seems to occur more frequently than exposure to high or low levels of all three exposures. Hence, we estimated the association of a combination of increased air pollution and road traffic noise and decreased surrounding green. The joint odds ratio (JOR) represents the odds

**Table 1.** Characteristics of the population ( $n = 354,827$ ).

Characteristic	Category	<i>n</i> (%)
Sex	Female	193,782 (54.6)
Age (years)	19–39	68,940 (19.4)
	40–64	134,161 (37.8)
	≥ 65	151,726 (42.8)
	Married, living together	246,775 (70.6)
Marital status	Unmarried/never married	42,233 (12.1)
	Divorced	22,509 (6.4)
	Widowed	38,089 (10.9)
	Dutch	308,167 (86.8)
Country of origin	Other, Western <sup>a</sup>	29,215 (8.2)
	Other, non-Western <sup>b</sup>	6,543 (1.8)
	Netherlands Antilles	1,500 (0.4)
	Suriname	3,820 (1.1)
	Turkey	3,284 (0.9)
	Morocco	2,298 (0.6)
Education	Primary or less	33,679 (9.8)
	Lower secondary	119,549 (34.9)
	Higher secondary	97,005 (28.3)
	University	92,733 (27.0)
Work	Yes	158,704 (48.1)
Standardized household income <sup>c</sup>	<€15,200	35,539 (10.1)
	€15,200–19,399	67,391 (19.1)
	€19,400–24,199	74,677 (21.2)
	€24,200–30,999	83,916 (23.8)
	≥€31,000	91,224 (25.9)
Smoking habits	Current	65,702 (19.7)
	Former	133,267 (39.9)
	Never	134,794 (40.4)
Number of cigarettes/day for current smokers [mean (SD)]	—	10.2 (8.3)
Alcohol consumption	Current	280,311 (82.2)
	Former	20,564 (6.0)
	Never	40,007 (11.7)
Number of alcohol glasses/week for current consumers [mean (SD)]	—	8.5 (9.6)
BMI	<18.5 kg/m <sup>2</sup>	4,628 (1.3)
	18.5–24.9 kg/m <sup>2</sup>	154,595 (44.3)
	25.0–30.0 kg/m <sup>2</sup>	130,058 (37.2)
	>30.0 kg/m <sup>2</sup>	59,987 (17.2)
	≤340 min/wk	83,340 (25.1)
Physical activity	340–720 min/wk	86,401 (26.0)
	720–1,380 min/wk	80,470 (24.2)
	>1,380 min/wk	81,634 (24.6)
	≤30	75,223 (21.3)
SES neighborhood <sup>d</sup>	30–34	78,723 (22.3)
	35–38	73,947 (20.9)
	38–43	58,521 (16.5)
	>43	67,342 (19.0)
	Yes	28,312 (8.0)
Self-reported physician-diagnosed diabetes	Yes	65,407 (21.6)
Self-reported physician-diagnosed hypertension	Yes	6,975 (2.1)
Self-reported physician-diagnosed stroke	Yes	11,866 (3.5)
Self-reported physician-diagnosed heart attack	Yes	11,866 (3.5)
NDVI 300 m [median (IQR)]	—	0.52 (0.13)
NDVI 1,000 m [median (IQR)]	—	0.56 (0.14)
TOP10NL 300 m [median (IQR)] <sup>e</sup>	—	0.24 (0.24)
TOP10NL 1,000 m [median (IQR)] <sup>f</sup>	—	0.40 (0.32)
PM <sub>10</sub> [median (IQR)]	μg/m <sup>3</sup>	24.4 (1.24)
PM <sub>coarse</sub> [median (IQR)]	μg/m <sup>3</sup>	8.1 (0.82)
PM <sub>2.5</sub> [median (IQR)]	μg/m <sup>3</sup>	16.7 (0.83)
PM <sub>2.5abs</sub> [median (IQR)]	10 <sup>–5</sup> /m	1.2 (0.25)
NO <sub>2</sub> [median (IQR)] <sup>g</sup>	μg/m <sup>3</sup>	23.4 (7.85)
OP <sup>DTT</sup> [median (IQR)]	nmol DTT/min/m <sup>3</sup>	1.19 (0.27)
OP <sup>ESR</sup> [median (IQR)]	A.U./1,000/m <sup>3</sup>	0.89 (0.18)
Road traffic noise [median (IQR)]	Lden (dB)	53.3 (7.5)

Note: —, no data; BMI, body mass index; CI, confidence interval, IQR, interquartile range, NDVI, Normalized Difference Vegetation Index; NO<sub>2</sub>, nitrogen dioxide; OP<sup>DTT</sup>, oxidative potential (OP) metric with dithiothreitol (DTT); OP<sup>ESR</sup>, oxidative potential (OP) metric with electron spin resonance; PM<sub>2.5</sub>, PM with aerodynamic diameter ≤2.5 μm; PM<sub>2.5abs</sub>, black carbon (PM<sub>2.5</sub> absorbance); PM<sub>10</sub>, particulate matter with aerodynamic diameter ≤10 μm; SD, standard deviation; SES, socioeconomic status; A.U., arbitrary unit.

<sup>a</sup>Other Western: Europe, North America, Oceania, Indonesia, Japan.

<sup>b</sup>Other non-Western: Africa, Latin America, Asia (excluding Indonesia and Japan).

<sup>c</sup>Standardized household income is adjusted for differences in household size and composition.

<sup>d</sup>SES neighborhood represents the educational, occupational, and economical status of the neighborhood.

<sup>e</sup>Surrounding green space in a 300-m buffer based on TOP10NL.

<sup>f</sup>Surrounding green space in a 1,000-m buffer based on TOP10NL.

<sup>g</sup>NO<sub>2</sub> values higher than 80 μg/m<sup>3</sup> were set at 80 μg/m<sup>3</sup>.

for a 1-unit (IQR) increase in air pollution and road traffic noise and a 1-unit decrease in surrounding green exposure relative to the odds for no increase (no decrease in surrounding green) in any of the exposures. The direction of the associations of surrounding green in the JORs are opposite to the direction used to describe associations of surrounding green in single- and multi-exposure models.

We denote the JOR based on the combination of the  $p$  exposures evaluated at  $x$  as the CRI and define it as

$$CRI = \exp\left(\sum_{p=1}^p \hat{\beta}_p x_p\right) = \exp(\hat{\beta}'x) = \prod_{p=1}^p JOR_p$$

where  $\hat{\beta}' = (\hat{\beta}_1, \dots, \hat{\beta}_p)$  are the estimates of the log odds ratio (OR) for the  $p$  exposures estimated in a logistic regression model consisting of all  $p$  exposures together,  $x' = (x_1, \dots, x_p)$  are the levels at which each exposure-specific OR is evaluated, and  $JOR_p = \exp(\hat{\beta}_p x_p)$  denotes the JOR for the  $p^{th}$  exposure in a multi-exposure model. JORs were estimated assuming additive effect estimates (log ORs) of joint exposures. The 95% CI of CRI is

defined by:  $\exp\left(\hat{\beta}'x \pm 1.96 \times \sqrt{x' \times Cov(\hat{\beta}) \times x}\right)$ . This definition of the CI is similar to that described by Crouse et al. (2015) and Jerrett et al. (2013). Interactions are not formally incorporated in the CRI method, as only main terms of exposures are included in the regression model used to calculate the CRI.

**Interaction analyses.** To evaluate potential interactions of combined exposures, we specified interaction terms in two-exposure models. To be able to observe patterns of interactions, we assessed interactions by combining a continuous exposure with quintiles of another exposure and vice versa, and by continuous exposure terms. Interactions were assessed on the multiplicative scale, the extent to which, on the OR scale, the effect estimate of being exposed to both exposures together exceeds the product of the effect estimates of the two exposures considered separately. We hypothesized that the association with exposure to air pollution is strongest (increased odds) in the highest road traffic noise quintile and vice versa. Further, we hypothesized that the association with air pollution and road traffic noise is strongest (increased odds) in the lowest surrounding green quintile and that the association with surrounding green is strongest (decreased odds) in the lowest air pollution or road traffic noise quintile. The rationale for this is that there is some evidence that associations of air pollution are stronger for people with psychological distress (Hicken et al. 2014) and that people who live in greener areas generally report less psychological distress (Gascon et al. 2015).

**Mediation analyses.** We performed mediation analyses to evaluate whether decreased levels of air pollution and road traffic noise are possible mechanisms underlying potential beneficial associations of surrounding green on cardiometabolic outcomes. We think that the application of mediation analyses in our cross-sectional study is reasonable because we assess mediation between environmental factors where the relationship is immediate. Mediation analysis does not add to the general limitation of our cross-sectional study. Previous mediation analyses in a cohort study have also applied green and air pollution from the same study period (James et al. 2016). Here, we specified air pollution and road traffic noise as mediators in the surrounding green–health association and thus assumed a causal effect between surrounding green, air pollution, and road traffic noise. Mediation analyses were only performed when surrounding green was significantly associated (decreased OR) with a cardiometabolic outcome in a single-exposure model. Of the potential mediator variables (air pollutants and road traffic noise), we only selected the exposures that were significantly

associated (increased ORs) with the cardiometabolic outcome for the mediation analyses.

We used the mediation package to estimate the direct, indirect, and total effects and the proportion mediated (Imai et al. 2010; Tingley et al. 2014). The mediation package uses the counterfactual framework to clarify the assumptions needed for causal mediation. Briefly, we specified mediator models (mediator = surrounding green + covariates) and outcome models (outcome = surrounding green + mediator + covariates). Both models were adjusted for SES indicators and lifestyle factors (Model 4). Next, we included the model objects (mediator model and outcome model) in the mediate function and specified the “treat” (surrounding green) and “mediator” variables. We used continuous terms for the treat and mediator variables in the mediation analyses. Exposure–response curves of associations between surrounding green and the mediators showed small deviations from linearity. We assessed CIs by 300 nonparametric bootstrap simulations. Due to computational limitations, we were not able to run the recommended 1,000 bootstrap simulations (Tingley et al. 2014).

Four important underlying confounding assumptions are required for the mediation analyses: no unmeasured exposure–outcome confounding, no unmeasured exposure–mediator confounding, no unmeasured mediator–outcome confounding, and no mediator–outcome confounder is affected by the exposure (VanderWeele 2016). While we were not able to verify these assumptions, we included major confounders in our mediation analyses and hence believe these assumptions are reasonable.

## Results

### Public Health Monitor Statistics

Our study population consisted of 354,827 persons aged 19 y or older (Table 1). Due to oversampling of the elderly, almost 43% of the subjects were 65 y or older. Further, people of Dutch origin (86.8% compared with 78% in the general population) were over-represented in the PHM. More information on the study population can be found elsewhere (Statistics Netherlands 2015).

Of our study population, 8.0% reported physician-diagnosed diabetes. Treatment in the past 12 months by a physician for hypertension was reported by 21.6% of the population, and treatment by a physician for heart attack and stroke was reported by 3.5% and 2.1% of the population, respectively. The degree of overlap between the cardiometabolic diseases was moderate to low. Of all subjects, 20.8% reported only one cardiometabolic disease. Two, three, and four cardiometabolic diseases were reported by 5.3%, 0.6%, and <0.1% of the subjects, respectively. Agreements between the outcomes were low; kappa values ranged from 0.05 to 0.18.

### Exposure Distribution and Correlations

The variability of TOP10NL surrounding green space [expressed as median (IQR)] was larger compared with NDVI surrounding greenness (Table 1). The variability of PM<sub>2.5</sub> and PM<sub>10</sub> was substantially lower than the variability of the other pollutants, particularly NO<sub>2</sub> and PM<sub>2.5abs</sub>. Road traffic noise varied less than NO<sub>2</sub> but more than PM<sub>2.5</sub> and PM<sub>10</sub> (Table 1).

Surrounding green, air pollution, and road traffic noise exposures at the study participants' home addresses were overall moderately correlated. In general, surrounding green (both NDVI and TOP10NL) in a 1,000-m buffer was slightly more strongly negatively correlated with concentrations of air pollutants and road traffic noise compared with surrounding green in a 300-m buffer (Figure 2). There was no correlation between NDVI surrounding greenness and PM<sub>2.5</sub> ( $r_s \leq 0.05$ ). Road traffic noise correlated weak to

moderately with concentrations of air pollutants ( $0.20 < r_s < 0.50$ , Figure 2). Surrounding green was weakly positively correlated with neighborhood SES (more green in higher SES neighborhoods), while air pollution and road traffic noise were weakly negatively correlated with neighborhood SES (Figure 2). Spatial collinearity was therefore not a concern in our analyses.

### Single-Exposure Regression Models

Analyses with surrounding green as continuous terms and quintiles of surrounding green indicated a significant inverse association between surrounding green and diabetes in our main models (Table 2 for continuous term, Table S2 for quintiles). We found a weak inverse association between surrounding green and the odds of hypertension, which was significant in the analysis with surrounding green as continuous term and for the fifth quintile (Tables 2 and S2). We found no association between surrounding green and the odds of stroke or heart attack (Tables 2 and S2).

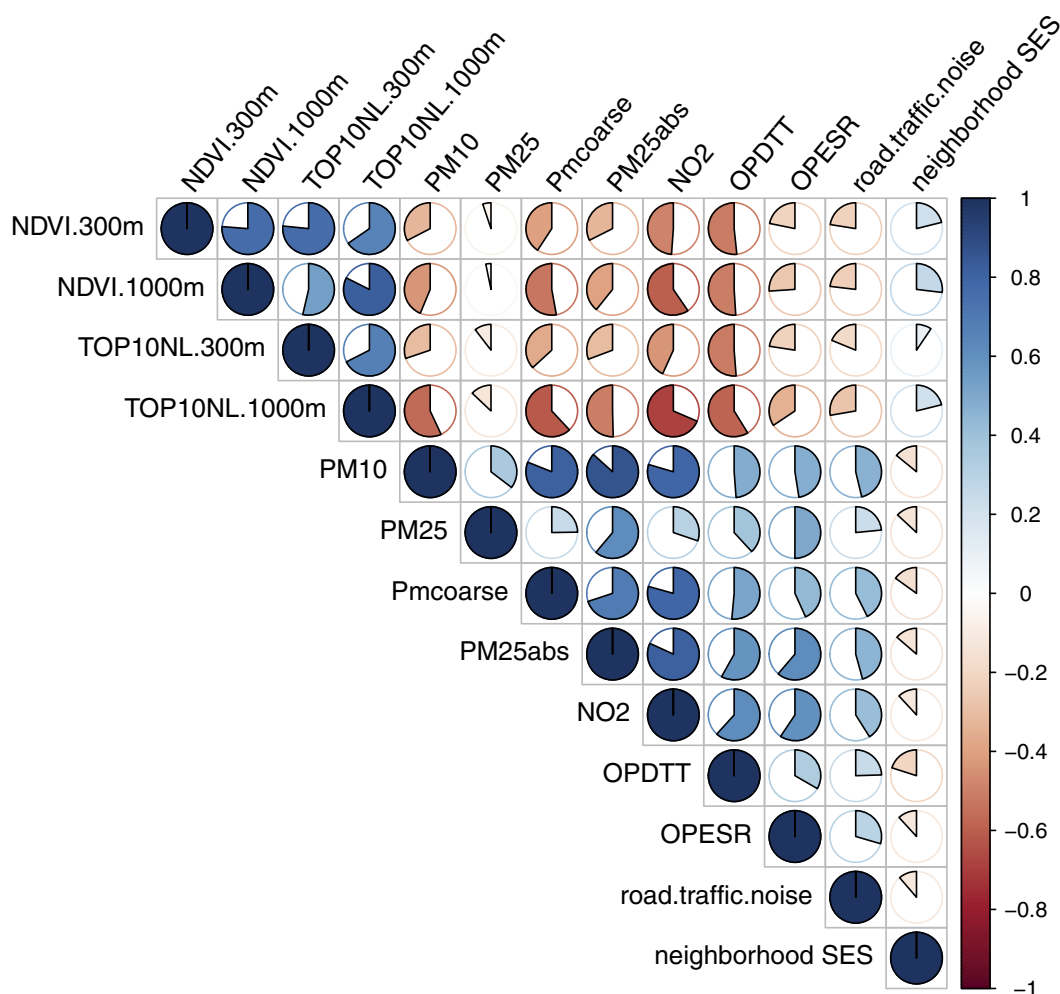
For most air pollutants, analyses with quintiles and with continuous terms showed a significant positive association with diabetes, especially for  $\text{NO}_2$  and  $\text{OP}^{\text{DTT}}$  (Tables 2 and S2).  $\text{NO}_2$ ,  $\text{OP}^{\text{DTT}}$ ,  $\text{OP}^{\text{ESR}}$ , and  $\text{PM}_{2.5\text{abs}}$  were significantly associated with higher odds of hypertension (Tables 2 and S2). ORs for stroke

and heart attack were increased for  $\text{PM}_{2.5}$  but not for other air pollutants (Tables 2 and S2).

Road traffic noise was positively associated with the odds of diabetes [ $\text{OR} = 1.02$ ; 95% CI: 1.00, 1.03] per 5 dB], but not with the odds of hypertension, stroke, or heart attack (Tables 2 and S2). Estimated exposure–response curves for diabetes and hypertension are shown in Figure S1.

Associations of surrounding green, air pollution, and road traffic noise with cardiometabolic outcomes were affected by the adjustment for covariates. In crude models, surrounding green, air pollution, and road traffic noise were significantly associated with diabetes, hypertension, stroke, and heart attack morbidity (Figure S2). Adjustment for individual SES and lifestyle factors attenuated the associations substantially ( $\geq 20\%$ ).

Sensitivity analyses showed similar associations with diabetes and hypertension in the elderly population ( $\geq 75$  y) compared with the full population. Associations of  $\text{PM}_{2.5}$  with stroke and heart attack, on the other hand, were somewhat stronger for the elderly population compared with the full population (Table S3). In analyses with heart attack ever and stroke ever, no significant associations were found with surrounding green or road traffic noise (Table S4). For air pollution, only  $\text{PM}_{2.5}$  was significantly associated with stroke ever ( $\text{OR} = 1.03$ ; 95% CI: 1.01, 1.06 per



**Figure 2.** Spearman correlations between surrounding green, air pollution, road traffic noise, and neighborhood socioeconomic status (SES). Clockwise pies denote positive correlations, and counterclockwise pies denote negative correlations. Note: All correlations are statistically significant, different from 0.0 ( $p < 0.005$ ). Surrounding green space in a 300-m buffer based on TOP10NL; surrounding green space in a 1,000-m buffer based on TOP10NL. NDVI, Normalized Difference Vegetation Index;  $\text{NO}_2$ , nitrogen dioxide;  $\text{OP}^{\text{DTT}}$ , oxidative potential (OP) metric with dithiothreitol (DTT);  $\text{OP}^{\text{ESR}}$ , oxidative potential (OP) metric with electron spin resonance;  $\text{PM}_{2.5}$ , PM with aerodynamic diameter  $\leq 2.5$   $\mu\text{m}$ ;  $\text{PM}_{10}$ , particulate matter with aerodynamic diameter  $\leq 10$   $\mu\text{m}$ .

**Table 2.** Results of single-exposure regression models.

Exposure variable (increment)	OR (95% CI)			
	Diabetes	Hypertension	Stroke	Heart attack
NDVI 300 m (0.13)	0.91 (0.89, 0.93)	0.97 (0.96, 0.99)	1.00 (0.96, 1.04)	1.00 (0.97, 1.03)
NDVI 1,000 m (0.14)	0.92 (0.90, 0.94)	0.99 (0.97, 1.01)	1.00 (0.96, 1.04)	1.04 (1.00, 1.07)
TOP10NL 300 m (0.24) <sup>a</sup>	0.93 (0.91, 0.95)	0.97 (0.95, 0.98)	1.00 (0.96, 1.04)	0.97 (0.94, 1.00)
TOP10NL 1,000 m (0.32) <sup>b</sup>	0.91 (0.88, 0.93)	0.98 (0.96, 0.99)	0.98 (0.93, 1.03)	0.99 (0.95, 1.03)
NO <sub>2</sub> (7.85)	1.06 (1.04, 1.09)	1.02 (1.00, 1.03)	1.00 (0.96, 1.04)	0.98 (0.95, 1.02)
OP <sup>DTT</sup> (0.27)	1.09 (1.07, 1.12)	1.06 (1.04, 1.07)	1.02 (0.97, 1.06)	1.03 (0.99, 1.06)
OP <sup>ESR</sup> (0.18)	1.02 (1.00, 1.04)	1.01 (1.00, 1.03)	1.02 (0.99, 1.06)	1.00 (0.97, 1.02)
PM <sub>10</sub> (1.24)	1.03 (1.01, 1.05)	1.00 (0.99, 1.01)	1.00 (0.96, 1.03)	0.97 (0.95, 1.00)
PM <sub>2.5</sub> (0.83)	1.01 (0.99, 1.03)	1.01 (1.00, 1.02)	1.05 (1.01, 1.09)	1.07 (1.05, 1.10)
PM <sub>2.5abs</sub> (0.25)	1.03 (1.01, 1.05)	1.01 (1.00, 1.02)	1.02 (0.98, 1.05)	1.01 (0.99, 1.04)
PM <sub>coarse</sub> (0.82)	1.03 (1.01, 1.05)	1.00 (0.99, 1.01)	1.01 (0.98, 1.05)	0.96 (0.94, 0.99)
Road traffic noise, Lden (5)	1.02 (1.00, 1.03)	1.00 (0.99, 1.01)	1.01 (0.98, 1.03)	1.00 (0.98, 1.02)

Note: Results are presented as OR (95% CI) per interquartile range increase (surrounding green and air pollution) or per 5 dB (road traffic noise). Models were adjusted for sex, age, marital status, region of origin, education, work, standardized household income, smoking habits, number of cigarettes/day, alcohol consumption, number of alcohol glasses/week, physical activity, body mass index, and neighborhood socioeconomic status (SES). CI, confidence interval; Lden, daily average noise level; NDVI, Normalized Difference Vegetation Index; NO<sub>2</sub>, nitrogen dioxide; OP<sup>DTT</sup>, oxidative potential (OP) metric with dithiothreitol (DTT) assay; OP<sup>ESR</sup>, oxidative potential (OP) metric with electron spin resonance; OR, odds ratio; PM<sub>2.5</sub>, PM with aerodynamic diameter ≤ 2.5 μm; PM<sub>2.5abs</sub>, black carbon (PM<sub>2.5absorbance</sub>); PM<sub>10</sub>, particulate matter with aerodynamic diameter ≤ 10 μm.

<sup>a</sup>Surrounding green space in a 300-m buffer based on TOP10NL.

<sup>b</sup>Surrounding green space in a 1,000-m buffer based on TOP10NL.

IQR increase) and heart attack ever (OR = 1.04; 95% CI: 1.01, 1.06 per IQR increase), similar to the analyses of our main (doctor-diagnosed) end points. Associations of NO<sub>2</sub> were similar to associations of NO<sub>2</sub> when values above 80 μg/m<sup>3</sup> were not set to 80 μg/m<sup>3</sup> (Table S5).

### Multi-exposure Regression Models

Based on the findings of the single-exposure models, we decided to focus on diabetes and hypertension in multi-exposure models. We only used NDVI surrounding greenness exposure variables, as this is a commonly used green exposure measure in epidemiological studies. For air pollution, we used OP<sup>DTT</sup>, NO<sub>2</sub>, and PM<sub>2.5</sub>, as these variables were most strongly associated with at least one cardiometabolic disease among the air pollutants.

**Potential for confounding.** The association of NDVI surrounding greenness with diabetes attenuated slightly when adjusted for OP<sup>DTT</sup> and vice versa (Table 3 for a 300-m buffer; Table S6 for a 1,000-m buffer). The effect estimate (OR) of NDVI 300 m attenuated from 0.91 (95% CI: 0.89, 0.93) per IQR increase to 0.93 (95% CI: 0.91, 0.96) and the effect estimate (OR) of OP<sup>DTT</sup> attenuated from 1.09 (95% CI: 1.07, 1.12) per IQR increase to 1.05 (95% CI: 1.03, 1.08). Adjustment for road traffic noise hardly affected the associations of NDVI 300 m or OP<sup>DTT</sup> with diabetes. Associations of road traffic noise with diabetes attenuated and lost significance when adjusted for NDVI of 300 m or OP<sup>DTT</sup>.

Associations of OP<sup>DTT</sup> with hypertension slightly attenuated after adjustment for NDVI 300 m (Table 4). The weak associations of NDVI 300 m with hypertension became insignificant

**Table 3.** Associations of surrounding green (NDVI 300 m), air pollution (OP<sup>DTT</sup>), and road traffic noise (Lden) with diabetes from single- and two-exposure models.

Exposure variable	Exposure metric <sup>a</sup>	Two exposure (adj. for) <sup>b</sup>			
		OR (95% CI)			
		Single exposure	NDVI 300 m	OP <sup>DTT</sup>	Road traffic noise (Lden)
NDVI 300 m	Continuous (per IQR)	0.91 (0.89, 0.93)	—	0.93 (0.91, 0.96)	0.91 (0.89, 0.93)
	Q1 (≤ 0.44)	1.00	—	1.00	1.00
	Q2 (≤ 0.50)	0.98 (0.93, 1.03)	—	0.99 (0.94, 1.04)	0.98 (0.93, 1.03)
	Q3 (≤ 0.55)	0.97 (0.93, 1.02)	—	0.99 (0.94, 1.04)	0.97 (0.93, 1.02)
	Q4 (≤ 0.61)	0.91 (0.87, 0.96)	—	0.95 (0.90, 0.99)	0.92 (0.87, 0.96)
	Q5 (> 0.61)	0.80 (0.76, 0.84)	—	0.84 (0.79, 0.89)	0.80 (0.76, 0.85)
OP <sup>DTT</sup>	Continuous (per IQR)	1.09 (1.07, 1.12)	1.05 (1.03, 1.08)	—	1.09 (1.07, 1.12)
	Q1 (< 1.01)	1.00	1.00	—	1.00
	Q2 (≤ 1.14)	1.04 (0.99, 1.09)	0.99 (0.94, 1.05)	—	1.04 (0.98, 1.09)
	Q3 (≤ 1.24)	1.07 (1.02, 1.13)	1.00 (0.95, 1.06)	—	1.07 (1.01, 1.12)
	Q4 (≤ 1.35)	1.17 (1.11, 1.23)	1.08 (1.02, 1.14)	—	1.16 (1.11, 1.22)
	Q5 (> 1.35)	1.20 (1.14, 1.26)	1.10 (1.04, 1.16)	—	1.19 (1.13, 1.25)
Road traffic noise (Lden)	Continuous (per 5 dB)	1.02 (1.00, 1.03)	1.00 (0.99, 1.02)	1.00 (0.99, 1.02)	—
	Q1 (< 49.3)	1.00	1.00	1.00	—
	Q2 (≤ 52.0)	1.05 (1.00, 1.11)	1.03 (0.98, 1.08)	1.04 (0.98, 1.09)	—
	Q3 (≤ 54.7)	1.07 (1.01, 1.12)	1.03 (0.98, 1.09)	1.04 (0.99, 1.09)	—
	Q4 (≤ 58.7)	1.08 (1.03, 1.13)	1.04 (0.99, 1.10)	1.05 (1.00, 1.10)	—
	Q5 (> 58.7)	1.06 (1.01, 1.12)	1.03 (0.98, 1.08)	1.02 (0.97, 1.08)	—

Note: Results are presented as OR (95% CI) relative to the reference category for the quintile analyses (OR = 1.00) or per continuous increase in the linear model. Models were adjusted for sex, age, marital status, region of origin, education, work, standardized household income, smoking habits, number of cigarettes/day, alcohol consumption, number of alcohol glasses/week, physical activity, body mass index, and neighborhood socioeconomic status (SES). —, no data; CI, confidence interval; IQR, interquartile range; Lden, daily average noise level; NDVI, Normalized Difference Vegetation Index; OP<sup>DTT</sup>, oxidative potential (OP) metric with dithiothreitol (DTT) assay; OR, odds ratio.

<sup>a</sup>IQR for NDVI 300 m: 0.13, OP<sup>DTT</sup>: 0.27 nmol DTT/min/m<sup>3</sup>.

<sup>b</sup>In all two-exposure models, generalized variance inflation factor (GVIF) values were below 1.8.

**Table 4.** Associations of surrounding green (NDVI 300 m), air pollution (OP<sup>DTT</sup>), and road traffic noise (Lden) with hypertension in single- and two-exposure models.

Exposure variable	Exposure metric <sup>a</sup>	Two exposure (adj. for) <sup>b</sup>			
		OR (95% CI)			
		Single exposure	NDVI 300 m	OP <sup>DTT</sup>	Road traffic noise (Lden)
NDVI 300 m	Continuous (per IQR)	0.97 (0.96, 0.99)	—	0.99 (0.98, 1.01)	0.97 (0.96, 0.98)
	Q1 (≤0.44)	1.00	—	1.00	1.00
	Q2 (≤0.50)	1.03 (1.00, 1.07)	—	1.04 (1.00, 1.07)	1.03 (0.99, 1.07)
	Q3 (≤0.55)	1.02 (0.99, 1.06)	—	1.03 (1.00, 1.07)	1.02 (0.99, 1.06)
	Q4 (≤0.61)	1.00 (0.96, 1.03)	—	1.02 (0.98, 1.05)	0.99 (0.96, 1.03)
	Q5 (>0.61)	0.95 (0.91, 0.98)	—	0.99 (0.95, 1.03)	0.94 (0.91, 0.98)
OP <sup>DTT</sup>	Continuous (per IQR)	1.06 (1.04, 1.07)	1.05 (1.04, 1.07)	—	1.06 (1.04, 1.08)
	Q1 (<1.01)	1.00	1.00	—	1.00
	Q2 (≤1.14)	1.07 (1.03, 1.10)	1.06 (1.02, 1.09)	—	1.07 (1.03, 1.10)
	Q3 (≤1.24)	1.06 (1.02, 1.09)	1.04 (1.00, 1.08)	—	1.06 (1.02, 1.09)
	Q4 (≤1.35)	1.10 (1.07, 1.14)	1.08 (1.04, 1.13)	—	1.11 (1.07, 1.15)
	Q5 (>1.35)	1.12 (1.08, 1.16)	1.10 (1.06, 1.14)	—	1.12 (1.08, 1.16)
Road traffic noise (Lden)	Continuous (per 5 dB)	1.00 (0.99, 1.01)	0.99 (0.98, 1.00)	0.99 (0.98, 1.00)	—
	Q1 (<49.3)	1.00	1.00	1.00	—
	Q2 (≤52.0)	1.04 (1.01, 1.08)	1.03 (1.00, 1.07)	1.03 (1.00, 1.07)	—
	Q3 (≤54.7)	1.00 (0.97, 1.03)	0.99 (0.95, 1.02)	0.98 (0.95, 1.02)	—
	Q4 (≤58.7)	1.00 (0.97, 1.04)	0.99 (0.96, 1.03)	0.99 (0.95, 1.02)	—
	Q5 (>58.7)	1.00 (0.96, 1.03)	0.99 (0.95, 1.02)	0.98 (0.94, 1.01)	—

Note: Results are presented as OR (95% CI) relative to the reference category (OR = 1.00) or per continuous increase. Models were adjusted for sex, age, marital status, region of origin, education, work, standardized household income, smoking habits, number of cigarettes/day, alcohol consumption, number of alcohol glasses/week, physical activity, body mass index, and neighborhood socioeconomic status (SES). —, no data; CI, confidence interval; IQR, interquartile range; Lden, daily average noise level; NDVI, Normalized Difference Vegetation Index; OP<sup>DTT</sup>, oxidative potential (OP) metric with dithiothreitol (DTT) assay; OR, odds ratio.

<sup>a</sup>IQR for NDVI 300 m: 0.13, OP<sup>DTT</sup>: 0.27 nmol DTT/min/m<sup>3</sup>.

<sup>b</sup>In all two-exposure models, generalized variance inflation factor (GVIF) values were below 1.8.

when adjusted for OP<sup>DTT</sup>; the OR decreased from 0.97 (95% CI: 0.96, 0.99) to 0.99 (95% CI: 0.98, 1.01). Adjustments for road traffic noise did not affect the associations of OP<sup>DTT</sup> or NDVI 300 m with hypertension.

Associations of PM<sub>2.5</sub> with stroke and heart attack were not affected by adjustment for surrounding green and traffic noise (Table S6), consistent with the lack of associations of these exposures with stroke and heart attack.

**Joint odds ratio.** For diabetes, we found the largest JOR for a combination of OP<sup>DTT</sup> and decreased NDVI 300 m (JOR: 1.13; 95% CI: 1.10, 1.16). The combined JOR was larger than the single-exposure ORs and similar to the JOR of OP<sup>DTT</sup>, road traffic noise, and decreased NDVI 300 m and of NO<sub>2</sub>, OP<sup>DTT</sup>, road traffic noise, and decreased NDVI 300 m (Figure 3). For hypertension, the largest JOR was also found for a combination of OP<sup>DTT</sup> and decreased NDVI 300 m (JOR: 1.06; 95% CI: 1.04, 1.08). However, the JOR was similar to the OR of OP<sup>DTT</sup> in the single-exposure model (OR: 1.06; 95% CI: 1.04, 1.07).

**Potential for interaction.** We hypothesized that the association with exposure to air pollution is strongest (increased odds) in the highest road traffic noise quintile and vice versa. Further, we hypothesized that the association with air pollution and road traffic noise is strongest (increased odds) in the lowest surrounding green quintile and that the association with surrounding green is strongest (decreased odds) in the lowest air pollution or road traffic noise quintile. However, we found no indications for multiplicative interactions in the hypothesized directions between combinations of exposure variables and cardiometabolic outcomes. Interaction terms were nonsignificant with the exception of the NDVI 300 m–OP<sup>DTT</sup> interaction ( $p < 0.005$ ). In the lowest OP<sup>DTT</sup> quintile, an IQR increase in NDVI 300 m was associated with lower odds of diabetes (OR: 0.89; 95% CI: 0.85, 0.93) (Table 5). However, in the fourth and fifth OP<sup>DTT</sup> quintile, an IQR increase in NDVI 300 m was associated with higher odds of diabetes, opposite to the hypothesized direction of the interaction. In the lowest road traffic noise quintile, an IQR increase in NDVI 300 m was also associated with lower levels of diabetes (OR:

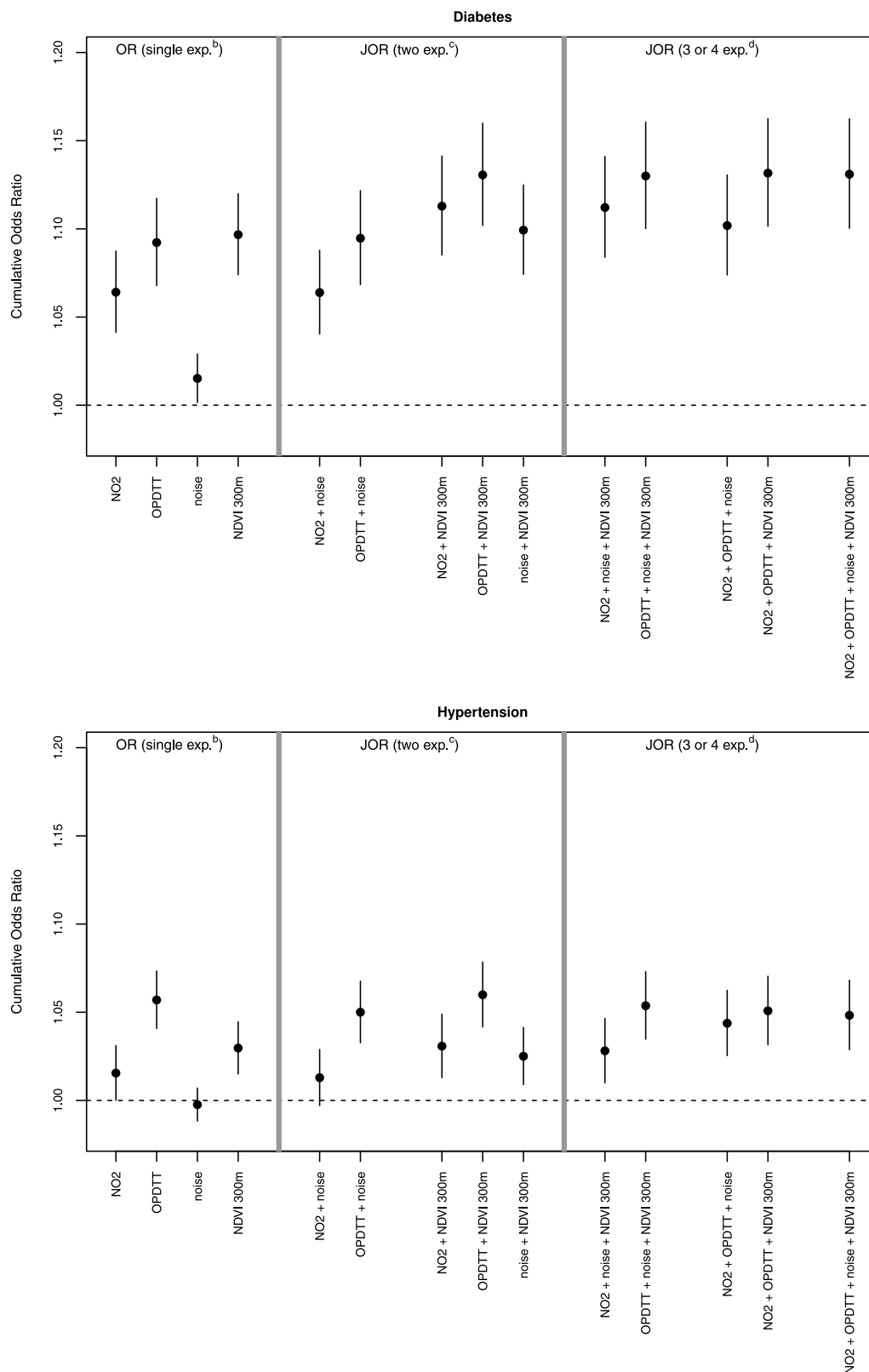
0.90; 95% CI: 0.85, 0.95). For the odds of hypertension, we also found no clear pattern (Table S7). Further, we also found no indications for multiplicative interactions in the hypothesized directions when we used continuous by continuous exposures terms (Table S8).

### Mediation Analyses

Based on the results of single-exposure regression models with NDVI surrounding greenness, we only conducted mediation analyses of surrounding greenness with the odds of diabetes. Assuming underlying assumptions of the mediation analyses hold, the association between surrounding greenness and diabetes can be partly explained by decreased levels of NO<sub>2</sub> or OP<sup>DTT</sup> (Table 6). The proportion mediated was higher for 1,000-m NDVI surrounding greenness compared with 300-m NDVI. Of the total estimated effect of 1,000-m NDVI, 38% (95% CI: 22, 60) could be explained by decreased OP<sup>DTT</sup> concentrations. For 300-m NDVI, the proportion mediated (via OP<sup>DTT</sup>) was 27% (95% CI: 13, 44).

### Discussion

In single-exposure models, we found that surrounding green was associated with decreased odds of diabetes, while exposure to air pollution and road traffic noise was associated with increased odds of diabetes. Road traffic noise was not associated with diabetes after adjustment for air pollution (OP<sup>DTT</sup>, NO<sub>2</sub>) or surrounding green, whereas associations with surrounding green and air pollution were attenuated but still evident in mutually adjusted models. JORs of combinations of air pollution and decreased surrounding green were higher than the ORs from the single-exposure models for diabetes. Air pollution and surrounding green, but not road traffic noise, were associated with hypertension in single-exposure models. The weak inverse association of surrounding green with hypertension attenuated and lost significance when adjusted for air pollution. Only PM<sub>2.5</sub> was associated with stroke and heart attack.



**Figure 3.** Joint odds ratios (JORs) for associations of air pollution, road traffic noise, and decreased surrounding green with diabetes and hypertension. The JORs are based on the CRI methods and represents the odds for a 1-unit [interquartile range (IQR)] increase in air pollution and road traffic noise and a 1-unit decrease in surrounding green exposure relative to the odds for no increase (no decrease in surrounding green) in any of the exposures. Results are given as JOR [95% confidence interval (CI)] per continuous increase for air pollution and road traffic noise and decrease for surrounding green (IQR for Normalized Difference Vegetation Index (NDVI) 300 m: 0.13, IQR for OP<sup>DTT</sup>: 0.27 nmol DTT/min/m<sup>3</sup>, IQR for NO<sub>2</sub>: 7.85 µg/m<sup>3</sup>, increment for road traffic noise: 5 dB) (main model). <sup>b</sup>ORs are based on effect estimates of single-exposure models, <sup>c</sup>JORs are based on effect estimates of two-exposure models, <sup>d</sup>JORs are based on effect estimates of three- or four-exposure models. Note: NO<sub>2</sub>, nitrogen dioxide; OP<sup>DTT</sup>, oxidative potential (OP) with the dithiothreitol (DTT) assay.

**Table 5.** Multiplicative interactions of surrounding green (NDVI 300 m), air pollution (OP<sup>DTT</sup>), and road traffic noise (Lden) on the odds of diabetes.

Stratified exposure variable	Quintile	Linear exposure variable		
		OR (95% CI)		
		NDVI 300 m	OP <sup>DTT</sup>	Road traffic noise (Lden)
NDVI 300 m	Q1 (≤0.44)	—	1.00 (0.94, 1.05)	0.98 (0.95, 1.01)
	Q2 (≤0.50)	—	1.01 (0.93, 1.09)	1.03 (0.99, 1.08)
	Q3 (≤0.55)	—	1.08 (1.00, 1.17)	1.03 (0.99, 1.08)
	Q4 (≤0.61)	—	1.08 (1.00, 1.17)	1.04 (1.00, 1.09)
	Q5 (>0.61)	—	1.09 (1.01, 1.18)	1.02 (0.98, 1.06)
	p.int <sup>a</sup>	—	0.05	0.34
OP <sup>DTT</sup>	Q1 (<1.01)	0.89 (0.85, 0.93)	—	1.00 (0.97, 1.02)
	Q2 (≤1.14)	0.99 (0.92, 1.06)	—	1.00 (0.96, 1.05)
	Q3 (≤1.24)	1.02 (0.95, 1.10)	—	1.01 (0.97, 1.06)
	Q4 (≤1.35)	1.12 (1.04, 1.20)	—	1.03 (0.98, 1.07)
	Q5 (>1.35)	1.17 (1.09, 1.25)	—	1.00 (0.96, 1.05)
	p.int <sup>a</sup>	<0.005	—	0.78
Road traffic noise (Lden)	Q1 (<49.3)	0.90 (0.86, 0.95)	1.08 (1.03, 1.13)	—
	Q2 (≤52.0)	0.97 (0.91, 1.04)	1.01 (0.94, 1.08)	—
	Q3 (≤54.7)	1.02 (0.95, 1.09)	1.00 (0.94, 1.08)	—
	Q4 (≤58.7)	1.02 (0.96, 1.09)	1.00 (0.93, 1.07)	—
	Q5 (>58.7)	1.03 (0.97, 1.10)	1.04 (0.97, 1.11)	—
	p.int <sup>a</sup>	0.48	0.78	—

Note: Results of multiplicative interactions are given as OR (95% CI) per 1-unit increase (IQR for NDVI 300 m: 0.13, IQR for OP<sup>DTT</sup>: 0.27 nmol DTT/min/m<sup>3</sup>, 5 dB for Lden) in quintiles of the second variable. Models were adjusted for sex, age, marital status, region of origin, education, work, standardized household income, smoking habits, number of cigarettes/day, alcohol consumption, number of alcohol glasses/week, physical activity, body mass index, and neighborhood socioeconomic status (SES). —, no data; CI, confidence interval; Lden, daily average noise level; NDVI, Normalized Difference Vegetation Index; OP<sup>DTT</sup>, oxidative potential (OP) metric with dithiothreitol (DTT) assay; OR, odds ratio.

<sup>a</sup>p.int shows the *p*-value for the overall interaction.

### Relationship between Environmental Exposures

Exposures to surrounding green, air pollution, and road traffic noise were, overall, moderately correlated, suggesting that confounding between the three exposures is possible. Because correlations were moderate, multicollinearity of exposures was not an issue in our regression analyses. We were therefore able to study the individual and combined associations of exposure to surrounding green, air pollution, and road traffic noise with cardio-metabolic disease.

The moderate negative correlations of surrounding green with air pollutants and road traffic noise that we found were also reported by other studies (Hystad et al. 2014; Thiering et al. 2016; Dadvand et al. 2012). The relationship between surrounding green and air pollution is probably due to several mechanisms. A causal relationship between surrounding green and air pollution could be removal of air pollutants from the air by uptake via leaf stomata or deposition to leaf surfaces, especially of trees. However, according to several studies, this effect is likely small and mainly limited to PM<sub>10</sub> (Setälä et al. 2013; Tallis et al. 2011; Nowak et al. 2006; Escobedo and Nowak 2009). Moreover, trees could increase air pollution concentrations by reducing dispersion of air pollution, especially in urban canyons. Another more likely explanation for the negative correlation between surrounding green and air pollution is the absence of air pollution sources within green areas (Wesseling et al. 2011).

**Table 6.** Mediation analyses of NDVI surrounding greenness (binary) and mediators NO<sub>2</sub> and OP<sup>DTT</sup> (continuous) on the odds of diabetes.

Indicator	Mediator	Proportion mediated (95% CI)
NDVI 300 m	OP <sup>DTT</sup>	0.27 (0.13, 0.44)
NDVI 300 m	NO <sub>2</sub>	0.20 (0.08, 0.33)
NDVI 1,000 m	OP <sup>DTT</sup>	0.38 (0.22, 0.60)
NDVI 1,000 m	NO <sub>2</sub>	0.34 (0.15, 0.53)

Note: Models were adjusted for sex, age, marital status, region of origin, education, work, standardized household income, smoking habits, number of cigarettes/day, alcohol consumption, number of alcohol glasses/week, physical activity, body mass index, and neighborhood socioeconomic status (SES). CI, confidence interval; NDVI, Normalized Difference Vegetation Index; NO<sub>2</sub>, nitrogen dioxide; OP<sup>DTT</sup>, oxidative potential (OP) metric with dithiothreitol (DTT) assay.

Exposure to road traffic noise was moderately correlated (*rs* between 0.23 and 0.46) with air pollution exposure in our study. This is in line with other studies, according to a review by Tétéreault et al. (2013). Correlations with road traffic noise were higher for pollutants with strong local traffic sources (NO<sub>2</sub>, PM<sub>2.5abs</sub>) than for pollutants with lower impact of local traffic (PM<sub>2.5</sub>, OP<sup>DTT</sup>). The moderate correlation of road traffic noise and traffic-related air pollutants is likely due to different features of local traffic affecting noise and pollution (Boogaard et al. 2009). Both exposures increase with higher traffic intensity in a street. Lower traffic speed in urban areas is typically associated with lower noise levels but can result in higher air pollution emissions. Road surface will affect traffic noise but not air pollution.

In this study, air pollution and road traffic noise exposures were predicted by models, which are (partly) based on land-use variables. Due to the use of similar land-use predictor variables and models with modest explained variability, correlations between modeled exposures can be higher or lower than between true exposures, respectively. Also, in the LUR model of OP<sup>DTT</sup> and in the STAMINA model (road traffic noise) but not in other models, an indicator of green space was included as a (marginal) predictor (Yang et al. 2015; Schreurs et al. 2010), potentially leading to overestimation of the relationship between surrounding green and OP<sup>DTT</sup> and road traffic noise. However, the correlation coefficients we found between modeled air pollution and noise exposures are similar to studies that evaluated correlations between measured air pollution and noise exposures in cities (Davies et al. 2009; Allen et al. 2009). Moreover, correlations between measured air pollution concentrations at the Dutch ESCAPE measurement sites and surrounding green and modeled road traffic noise were similar to correlations of modeled air pollution concentrations (Table S9).

### Confounding

In our single-exposure models, surrounding green, air pollution, and road traffic noise were associated with diabetes. This is in line with other studies that evaluated exposures to these exposures (Brown et al. 2016; Astell-Burt et al. 2014; Dalton et al.

2016; Eze et al. 2015; Balti et al. 2014; Wang et al. 2014a; Dzhambov 2015).

We found that the association between air pollution and diabetes was confounded by surrounding green, and vice versa. The majority of the studies that evaluated the link between air pollution or surrounding green and diabetes did not adjust for the other exposure (Dalton et al. 2016; Astell-Burt et al. 2014; Ngom et al. 2016; Maas et al. 2009; Eze et al. 2015; Balti et al. 2014). Findings of other studies regarding mutual confounding of surrounding green, air pollution, and noise associations are inconsistent. Eze et al. (2014) reported that the association between NO<sub>2</sub> and diabetes was attenuated but still positive after adjustment for road traffic noise. However, road traffic noise was not associated with diabetes after adjustment for NO<sub>2</sub> (Eze et al. 2014). In another study, Eze et al. (2017) found that road traffic noise remained associated with diabetes after adjustment for NO<sub>2</sub> or surrounding green (which were both not associated with diabetes). Sørensen et al. (2013) showed that road traffic noise exposure remained associated with the odds of diabetes after adjustment for NO<sub>x</sub>.

The weak association between surrounding green and hypertension that we found in single-exposure models attenuated and lost significance after adjustment for air pollution, whereas adjustments for surrounding green or road traffic noise barely affected associations of air pollution with hypertension. The lack of an association with road traffic noise is in line with a recent meta-analysis (Fuks et al. 2017).

Results of the CRI confirmed the results of our two-exposure models for diabetes and hypertension. The JOR of exposure to a combination of air pollution and decreased surrounding green for diabetes is underestimated by ORs from single-exposure models. As the inclusion of NO<sub>2</sub> or road traffic noise does not increase the JOR of OP<sup>DTT</sup> and decreased NDVI surrounding greenness, a combination of OP<sup>DTT</sup> and decreased NDVI surrounding greenness is sufficient to characterize the JOR of combined exposures to air pollution, road traffic noise, and decreased surrounding greenness on diabetes. For hypertension, the CRI showed that the JOR of exposure to a combination of air pollution and decreased surrounding green or the JOR of exposure to air pollution, road traffic noise, and decreased surrounding green is similar to the OR of air pollution from a single-exposure model. Of all environmental exposures, only PM<sub>2.5</sub> was associated with stroke and heart attack. Associations of PM<sub>2.5</sub> with stroke and heart attack were stronger in the elderly population compared with the full population. This might reflect greater susceptibility and/or reduced exposure misclassification in the elderly, since they may spend more time at or near their residences than younger adults. We did not find any association of surrounding green or road traffic noise with heart attack or stroke. Other studies evaluating associations of surrounding green, air pollution, and road traffic noise with stroke and heart attack morbidity showed mixed results. Some studies found associations with surrounding green or road traffic noise, while others reported no relationship (Atkinson et al. 2013; Babisch 2014; Brown et al. 2016; Cesaroni et al. 2014; Maas et al. 2009; Pereira et al. 2012; Stafoggia et al. 2014; Vienneau et al. 2015). Possible explanations for the different findings are differences in study design, adjustments for covariates, definitions of health outcomes, study populations, and exposure assessment methods (especially for surrounding green).

### Interaction

We did not find indications for interactions in the hypothesized direction for any cardiometabolic outcome. Some interaction terms were significant; however, no clear pattern was observed. As Hicken et al. (2014) reported that stronger associations of air pollution for people with psychological distress and traffic noise

(positively) and surrounding green (inversely) are associated with psychological distress (Seidler et al. 2017; Gascon et al. 2015), we hypothesized that the association with exposure to air pollution is strongest (increased odds) in the highest road traffic noise quintile and vice versa, that the association with air pollution and road traffic noise is strongest (increased odds) in the lowest surrounding green quintile, and that the association with surrounding green is strongest (decreased odds) in the lowest air pollution or road traffic noise quintile. Our results are in line with other studies that also reported no indications for interactions between air pollution and road traffic noise (Bodin et al. 2016; Sørensen et al. 2013; Selander et al. 2009), except for Sørensen et al. (2014). Hence, the strength of the association of an environmental exposure is probably not dependent on the level of another environmental exposure.

### Mediation

Assuming underlying assumptions of the mediation analyses hold, the association between NDVI surrounding greenness and diabetes was partly mediated by decreased concentrations of air pollution. The proportion mediated was dependent on the buffer size of surrounding greenness and on the air pollution variable.

The choice to treat air pollution and road traffic noise as confounder or mediator of the surrounding green–health association affects the effect size one should use to evaluate potential health impacts of surrounding green. If air pollution is primarily a confounder, the adjusted slopes for surrounding green from Table 3 (ORs of two-exposure models) should be used. Otherwise, the slopes from Table 2 (ORs of single-exposure models) should be used, and we expect that about 20–38% of the association between surrounding green and diabetes may be related to reduced air pollution. Higher greenness may lead to lower road traffic noise and air pollution concentrations by limiting dispersion by green barriers or by scavenging of air pollution. These mechanisms only provide a partial explanation of the empirical correlation between greenness and air pollution/road traffic noise. A stronger component is that in a greener area, there are fewer sources of air pollution and road traffic noise, and, therefore, pollution is lower. This reflects a common source (or lack thereof) and thus not a causal pathway from green to health.

Surrounding green may affect health via several pathways other than reduced air pollution; it may stimulate physical activity (and thereby reduce overweight) and decrease stress. In our mediation analyses, we adjusted for physical activity and weight status. Yet an estimated direct effect of surrounding green on diabetes remained. This might be due to lower stress levels.

### Strengths and Limitations

Strengths of this study include the large population size (>350,000 subjects) with national coverage and the ability to adjust for SES and lifestyle factors. Furthermore, we were able to study confounding, interaction (on the multiplicative scale), and mediation of several environmental exposures on diabetes, hypertension, stroke, and heart attack morbidity. We used multiple indicators of surrounding green and air pollution exposure, and all indicators were calculated at the address level. Moreover, all environmental exposures were assessed between 2009 and 2011, i.e., relatively close before the time the survey was administered (2012). The use of recent health and exposure data increases the relevance for current policies.

This study also has limitations. The cross-sectional study is more prone to potential selection bias than longitudinal designs. We had no information about the year of onset of cardiometabolic diseases, and hence, we do not know whether exposures

precede the health outcome. The biologically most relevant time period for effects of exposure to the environmental exposures (surrounding green, air pollution, and road traffic noise) and incidence of cardiometabolic outcomes is not well known. For air pollution, the hypothesis is that average exposure in the past 2 y is especially important for cardiometabolic diseases (Schwartz et al. 2008; Puett et al. 2008). Although we characterized exposure by 2009–2010 estimates, studies have shown that the spatial variation of air pollution and traffic noise exposure levels remain stable over periods of about 10 y in Western countries (Eeftens et al. 2011; Fecht et al. 2016). Current exposure estimates hence represent past exposure contrasts. However, subjects may have moved after development of the disease. This residential mobility may lead to some likely nondifferential exposure misclassification and subsequently an underestimation of the associations with health outcomes. Moreover, we did not have time–activity data, which may also lead to exposure error and potentially an underestimation of the associations. Further, we evaluated prevalence of a disease instead of incidence. As a result of their diseases, subjects may have changed their lifestyles.

Another limitation is the use of self-reported health outcomes, which may cause some error due to under- or overreporting. Studies evaluating the validity of self-reported health outcomes showed moderate to very good agreements between questionnaire responses and medical records for diabetes, hypertension, and myocardial infarction; agreements for stroke varied per study (Okura et al. 2004; Schneider et al. 2012; Hansen et al. 2014; Muggah et al. 2013; Machón et al. 2013). Strak et al. (2017) previously reported that 84.9% of this study population with self-reported diabetes has also been prescribed diabetes medication, based on an external dataset. Moreover, they observed similar associations between air pollution and prevalence of diabetes when investigating self-reported physician-diagnosed diabetes and diabetes medication prescription as separate outcomes. Fuks et al. (2017) found that air pollution was associated with self-reported hypertension but not with measured hypertension. The more consistent associations between environmental exposures and diabetes compared with the cardiovascular outcomes could be due to better reporting of the disease. Another explanation is that the exposures could be more strongly associated with fatal cardiovascular events. Several studies found associations of air pollution with fatal heart attacks and strokes but not or less strong with nonfatal heart attacks or strokes (Miller et al. 2007; Sørensen et al. 2014; Puett et al. 2009; Rosenlund et al. 2006, 2008, 2009). The associations reported in this study might be less strong than associations with total (fatal + nonfatal) heart attacks and strokes.

In this study, we evaluated the association between modeled exposures to air pollution and road traffic noise and cardiometabolic morbidity. Exposure assessment methods to determine air pollution concentrations and noise levels at residential addresses are well developed and commonly evaluated (Wang et al. 2013, 2014b; Schreurs et al. 2010). However, exposure assessment models may differ in their accuracy to predict the exposure. Performance of the LUR models differed between air pollutants (Table S1) and may have affected associations. A moderate LUR model performance could lead to weaker associations. Further, if LUR models predict air pollution concentrations at residential addresses more precisely than the STAMINA model predicts road traffic noise at residential addresses, it could explain the more robust associations of air pollution exposures in this study. Because of the lack of noise measurements, we cannot test this hypothesis. Surrounding green is measured and not modeled, but a major issue is how well-measured surrounding green reflects biologically relevant green exposure.

We were not able to verify the four important underlying confounding assumptions required for the mediation analyses. However, we included major confounders in our mediation analyses and hence believe that the assumptions are reasonable. We performed a mediation analysis in a cross-sectional analysis. The general limitations of cross-sectional studies with regard to the temporality of exposure–response relationships do not apply to the mediation analysis because we assessed mediation between environmental factors where the relationship is immediate. This would be different when we had performed a mediation analysis with an intermediate health marker (e.g., blood pressure) where the relationship may not be immediate. Further, we did not test for interactions between surrounding green and air pollution in the mediation analyses, as we did not find indications for interpretable interactions in the hypothesized direction in the two-exposure models.

## Conclusion

In single-exposure models, we found that surrounding green was associated with decreased odds of diabetes, while exposure to air pollution and road traffic noise was associated with increased odds of diabetes. After adjustment for air pollution (OP<sup>DTT</sup>, NO<sub>2</sub>) or surrounding green, the association of road traffic noise with diabetes attenuated and lost significance. Two-exposure analyses showed that associations of surrounding green and air pollution with diabetes generally remained, even though they were attenuated. Our results suggest that studies including only one of the correlated exposures (surrounding green, air pollution, and road traffic noise) may overestimate the association of diabetes attributed to that exposure. The joint impact of exposure to a combination of surrounding green and air pollution may be underestimated by the associations from single-exposure models. Future studies would benefit from including multiple environmental exposures to evaluate individual and combined associations.

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